

## DISCRETE MEMORYLESS SOURCE (DMS) Review

- The source output is an unending sequence,  $X_1, X_2, X_3, \dots$ , of random letters, each from a finite alphabet  $\mathcal{X}$ .
- Each source output  $X_1, X_2, \dots$  is selected from  $\mathcal{X}$  using a common probability measure with pmf  $p_X(x)$ .
- Each source output  $X_k$  is statistically independent of all other source outputs  $X_1, \dots, X_{k-1}, X_{k+1}, \dots$ .
- Without loss of generality, let  $\mathcal{X}$  be  $\{1, \dots, M\}$  and denote  $p_X(i)$ ,  $1 \leq i \leq M$  as  $p_i$ .

**OBJECTIVE:** Minimize expected length  $\bar{L}$  of prefix codes for a given DMS.

Let  $l_1, \dots, l_M$  be integer codeword lengths.

$$\bar{L}_{\min} = \min_{l_1, \dots, l_M: \sum 2^{-l_i} \leq 1} \left\{ \sum_{i=1}^M p_i l_i \right\}$$

Without the integer constraint,  $l_i = -\log p_i$  minimizes  $\bar{L}_{\min}$ , so

$$l_i = -\log p_i \quad (\text{desired length})$$

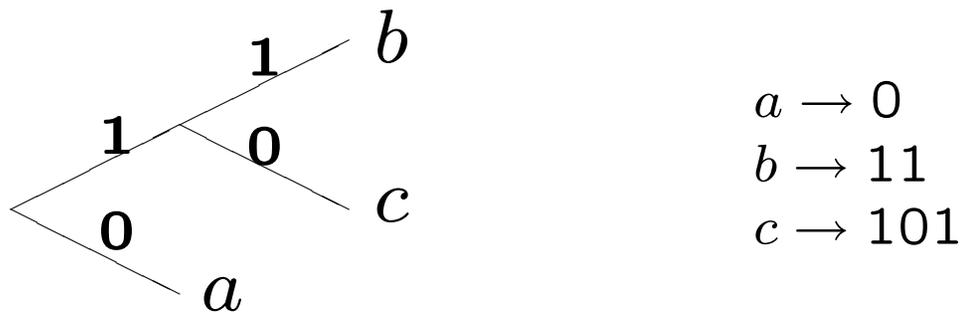
$$\bar{L}_{\min}(\text{non-int}) = \sum_i -p_i \log p_i \doteq \mathbf{H}(X)$$

$\mathbf{H}(X)$  is the entropy of  $X$ . It is the expected value of  $-\log p(X)$  and the desired expected length of the binary codeword.

**Theorem:** Let  $\bar{L}_{min}$  be the minimum expected codeword length over all prefix-free codes for  $X$ . Then

$$\mathbf{H}(X) \leq \bar{L}_{min} < \mathbf{H}(X) + 1$$

$\bar{L}_{min} = \mathbf{H}(X)$  iff each  $p_i$  is integer power of 2.

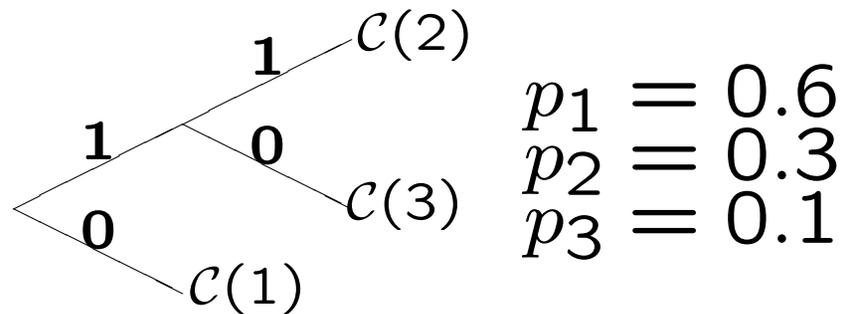


**Note that if  $p(a) = 1/2$ ,  $p(b) = 1/4$ ,  $p(c) = 1/4$ , then each binary digit is IID,  $1/2$ . This is general.**

## Huffman Coding Algorithm

Above theorem suggested that good codes have  $l_i \approx \log(1/p_i)$ .

Huffman took a different approach and looked at the tree for a prefix-free code.



**Lemma:** Optimal prefix-free codes have the property that if  $p_i > p_j$  then  $l_i \leq l_j$ . This means that  $p_i > p_j$  and  $l_i > l_j$  can't be optimal.

**Lemma:** Optimal prefix-free codes are full.

The sibling of a codeword is the string formed by changing the last bit of the codeword.

**Lemma:** For optimality, the sibling of each maximal length codeword is another codeword.

**Assume that**  $p_1 \geq p_2 \geq \dots \geq p_M$ .

**Lemma:** There is an optimal prefix-free code in which  $C(M-1)$  and  $C(M)$  are maximal length siblings.

Essentially, the codewords for  $M-1$  and  $M$  can be interchanged with max length codewords.

The Huffman algorithm first combines  $C(M-1)$  and  $C(M)$  and looks at the reduced tree with  $M-1$  nodes.

After combining two least likely codewords as siblings, we get a “reduced set” of probabilities.

symbol	$p_i$	
1	0.4	
2	0.2	
3	0.15	
4	0.15	1
5	0.1	0
		0.25

Finding the optimal code for the reduced set results in an optimal code for original set. Why?

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For any code for the reduced set  $X'$ , let expected length be  $\bar{L}'$ .

The expected length of the corresponding code for  $X$  has  $\bar{L} = \bar{L}' + p_{M-1} + p_M$ .

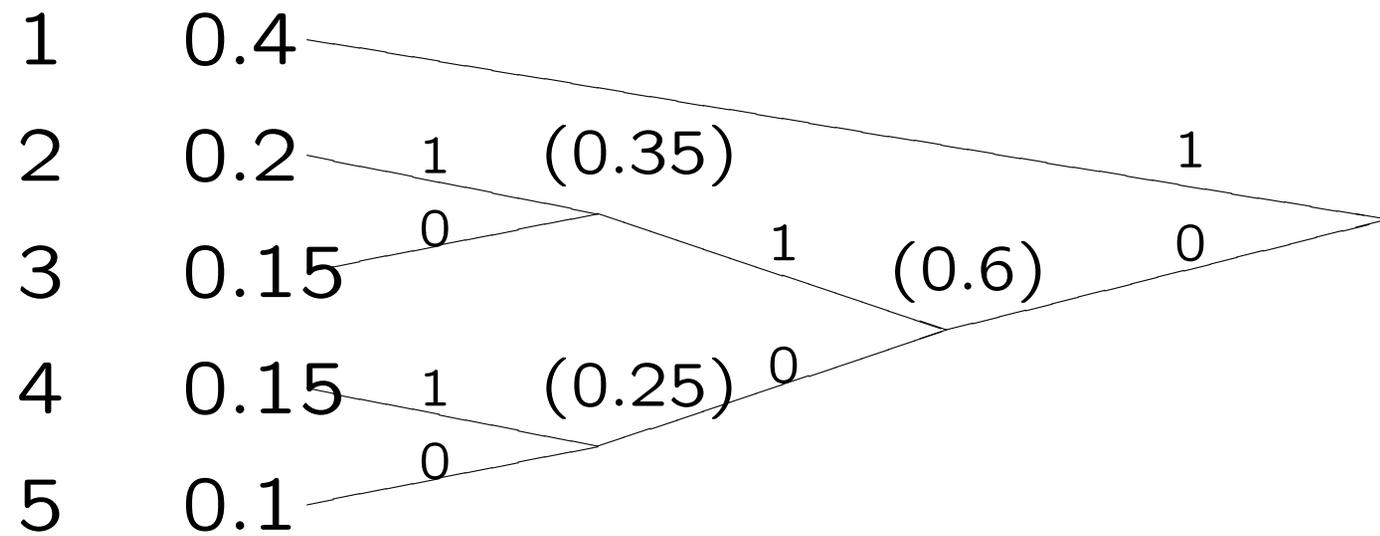
symbol	$p_i$	
1	0.4	
2	0.2	
3	0.15	
4	0.15	1
5	0.1	0
		0.25

Now we can tie together (siblingify?) the least two probable nodes in the reduced set.

symbol	$p_i$	
1	0.4	
2	0.2	
3	0.15	
4	0.15	1
5	0.1	0
		0.25

1	0.4		
2	0.2	1	
3	0.15	0	0.35
4	0.15	1	
5	0.1	0	
			0.25

Surely the rest is obvious.



## DISCRETE SOURCE CODING: REVIEW

The Kraft inequality,  $\sum_i 2^{-l_i} \leq 1$ , is a necessary and sufficient condition on prefix-free code-word lengths.

Given a pmf,  $p_1, \dots, p_M$  on a set of symbols, the Huffman algorithm constructs a prefix-free code of minimum expected length,  $\bar{L}_{\min} = \sum_i p_i l_i$ .

A discrete memoryless source (DMS) is a sequence of iid discrete chance variables  $X_1, X_2, \dots$ . The entropy of a DMS is  $H(X) = \sum_i -p_i \log(p_i)$ .

**Theorem:**  $H(X) \leq \bar{L}_{\min} < H(X) + 1$ .

**ENTROPY OF  $X$ ,  $|\mathcal{X}| = M$ ,  $\Pr(X=i) = p_i$**

$$\mathbf{H}(X) = \sum_i -p_i \log p_i = \mathbf{E}[-\log p_X(X)]$$

**$-\log p_X(X)$  is a rv, called the log pmf.**

**$\mathbf{H}(X) \geq 0$ ; Equality if  $X$  deterministic.**

**$\mathbf{H}(X) \leq \log M$ ; Equality if  $X$  equiprobable.**

**For independent rv's  $X, Y$ ,  $XY$  is also a chance variable taking on the sample value  $xy$  with probability  $p_{XY}(xy) = p_X(x)p_Y(y)$ .**

$$\begin{aligned} \mathbf{H}(XY) &= \mathbf{E}[-\log p(XY)] = \mathbf{E}[-\log p(X)p(Y)] \\ &= \mathbf{E}[-\log p(X) - \log p(Y)] = \mathbf{H}(X) + \mathbf{H}(Y) \end{aligned}$$

For a discrete memoryless source, a block of  $n$  random symbols,  $X_1, \dots, X_n$ , can be viewed as a single random symbol  $\mathbf{X}^n$  taking on the sample value  $\mathbf{x}^n = x_1 x_2 \dots x_n$  with probability

$$p_{\mathbf{X}^n}(\mathbf{x}^n) = \prod_{i=1}^n p_X(x_i)$$

The random symbol  $\mathbf{X}^n$  has the entropy

$$\begin{aligned} \mathbf{H}(\mathbf{X}^n) &= \mathbf{E}[-\log p(\mathbf{X}^n)] = \mathbf{E}\left[-\log \prod_{i=1}^n p_X(X_i)\right] \\ &= \mathbf{E}\left[\sum_{i=1}^n -\log p_X(X_i)\right] = n\mathbf{H}(X) \end{aligned}$$

## Fixed-to-variable prefix-free codes

Segment input into  $n$ -blocks  $X^n = X_1X_2\dots X_n$ .

Form min-length prefix-free code for  $X^n$ .

This is called an  $n$ -to-variable-length code

$$\mathbf{H}(X^n) = n\mathbf{H}(X)$$

$$\mathbf{H}(X^n) \leq \mathbf{E}[L(X^n)]_{\min} < \mathbf{H}(X^n) + 1$$

$$\bar{L}_{\min,n} = \frac{\mathbf{E}[L(X^n)]_{\min}}{n} \quad \text{bpss}$$

$$\mathbf{H}(X) \leq \bar{L}_{\min,n} < \mathbf{H}(X) + 1/n$$

$$\bar{L}_{\min,n} \rightarrow \mathbf{H}(X)$$

## WEAK LAW OF LARGE NUMBERS (WLLN)

Let  $Y_1, Y_2, \dots$  be sequence of rv's with mean  $\bar{Y}$  and variance  $\sigma_Y^2$ .

The sum  $A = Y_1 + \dots + Y_n$  has mean  $n\bar{Y}$  and variance  $n\sigma_Y^2$

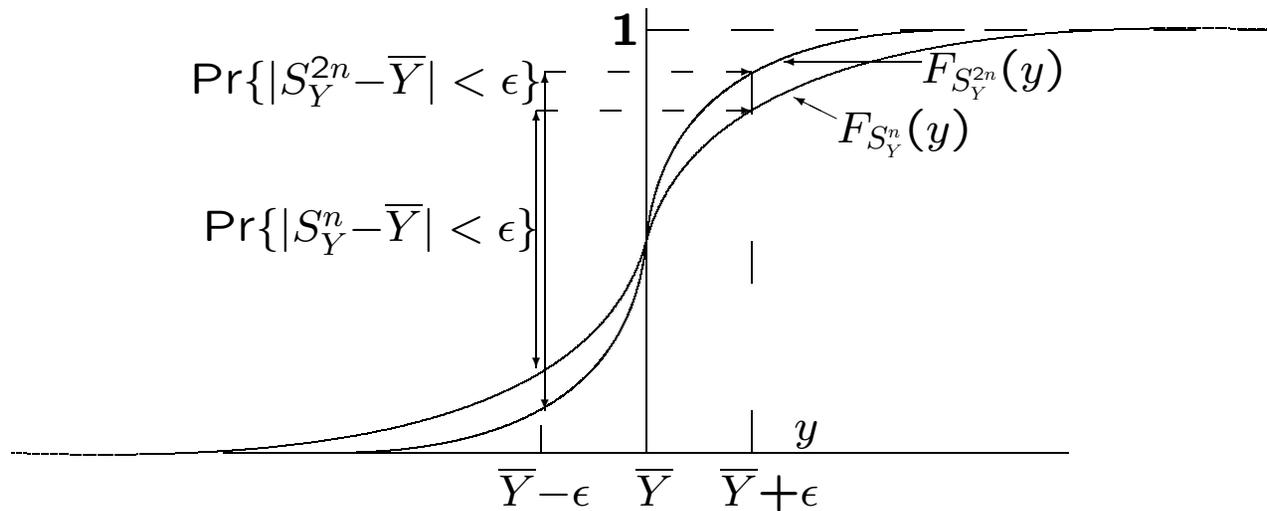
The sample average of  $Y_1, \dots, Y_n$  is

$$S_Y^n = \frac{A}{n} = \frac{Y_1 + \dots + Y_n}{n}$$

It has mean and variance

$$\mathbf{E}[S_Y^n] = \bar{Y}; \quad \mathbf{VAR}[S_Y^n] = \frac{\sigma_Y^2}{n}$$

**Note:**  $\lim_{n \rightarrow \infty} \mathbf{VAR}[A] = \infty$      $\lim_{n \rightarrow \infty} \mathbf{VAR}[S_Y^n] = 0$ .



The distribution of  $S_Y^n$  clusters around  $\bar{Y}$ , clustering more closely as  $n \rightarrow \infty$ .

**Chebyshev:** for  $\epsilon > 0$ ,  $\Pr\{|S_Y^n - \bar{Y}| \geq \epsilon\} \leq \frac{\sigma_Y^2}{n\epsilon^2}$

**For any  $\epsilon, \delta > 0$ , large enough  $n$ ,**

$$\Pr\{|S_Y^n - \bar{Y}| \geq \epsilon\} \leq \delta$$

## ASYMPTOTIC EQUIPARTITION PROPERTY (AEP)

Let  $X_1, X_2, \dots$ , be output from DMS.

Define log pmf as  $w(x) = -\log p_X(x)$ .

$w(x)$  maps source symbols into real numbers.

For each  $j$ ,  $W(X_j)$  is a rv; takes value  $w(x)$  for  $X_j = x$ . Note that

$$\mathbf{E}[W(X_j)] = \sum_x p_X(x) [-\log p_X(x)] = H(X)$$

$W(X_1), W(X_2), \dots$  sequence of iid rv's.

**For  $X_1 = x_1, X_2 = x_2$ , the outcome for  $W(X_1) + W(X_2)$  is**

$$\begin{aligned}w(x_1) + w(x_2) &= -\log p_X(x_1) - \log p_X(x_2) \\ &= -\log\{p_{X_1}(x_1)p_{X_2}(x_2)\} \\ &= -\log\{p_{X_1X_2}(x_1x_2)\} = w(x_1x_2)\end{aligned}$$

**where  $w(x_1x_2)$  is -log pmf of event  $X_1X_2 = x_1x_2$**

$$W(X_1X_2) = W(X_1) + W(X_2)$$

**$X_1X_2$  is a random symbol in its own right (takes values  $x_1x_2$ ).  $W(X_1X_2)$  is -log pmf of random symbol  $X_1X_2$ .**

**Probabilities multiply, log pmf's add.**

**For  $\mathbf{X}^n = \mathbf{x}^n$ ;  $\mathbf{x}^n = (x_1, \dots, x_n)$ , the outcome for  $W(X_1) + \dots + W(X_n)$  is**

$$\sum_{j=1}^n w(x_j) = - \sum_{j=1}^n \log p_X(x_j) = - \log p_{\mathbf{X}^n}(\mathbf{x}^n)$$

**Sample average of log pmf's is**

$$S_W^n = \frac{W(X_1) + \dots + W(X_n)}{n} = \frac{- \log p_{\mathbf{X}^n}(\mathbf{X}^n)}{n}$$

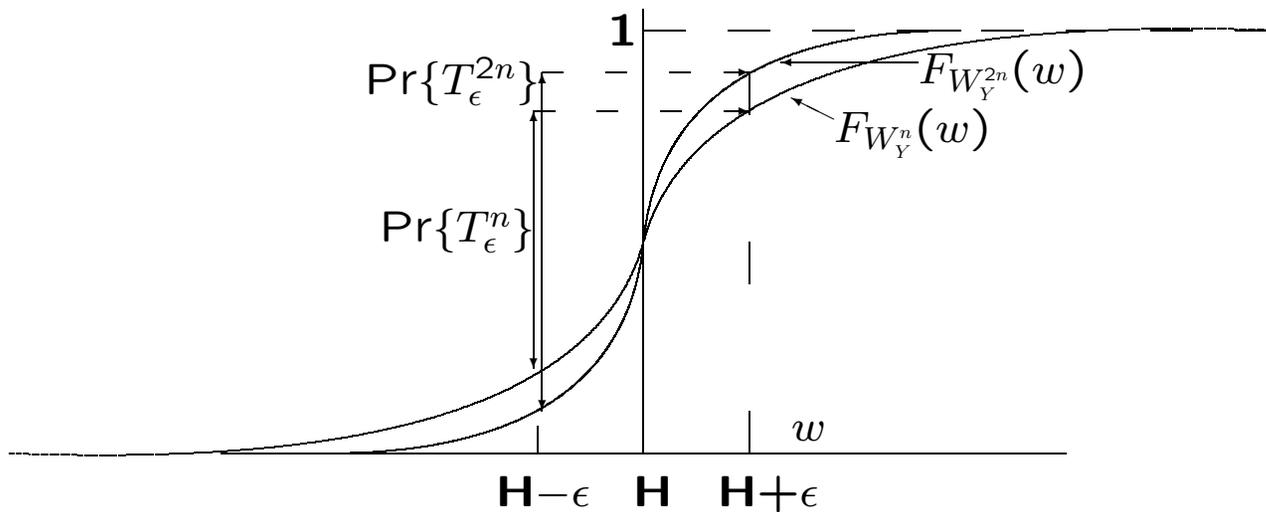
**WLLN applies and is**

$$\Pr \left( \left| S_W^n - \mathbf{E}[W(X)] \right| \geq \epsilon \right) \leq \frac{\sigma_W^2}{n\epsilon^2}$$

$$\Pr \left( \left| \frac{- \log p_{\mathbf{X}^n}(\mathbf{X}^n)}{n} - \mathbf{H}(X) \right| \geq \epsilon \right) \leq \frac{\sigma_W^2}{n\epsilon^2}.$$

Define typical set as

$$T_\epsilon^n = \left\{ \mathbf{x}^n : \left| \frac{-\log p_{\mathbf{X}^n}(\mathbf{x}^n)}{n} - \mathbf{H}(X) \right| < \epsilon \right\}$$



As  $n \rightarrow \infty$ , typical set approaches probability 1:

$$\Pr(\mathbf{X}^n \in T_\epsilon^n) \geq 1 - \frac{\sigma_W^2}{n\epsilon^2}$$

**We can also express  $T_\epsilon^n$  as**

$$T_\epsilon^n = \left\{ \mathbf{x}^n : n(\mathbf{H}(X) - \epsilon) < -\log p(\mathbf{x}^n) < n(\mathbf{H}(X) + \epsilon) \right\}$$

$$T_\epsilon^n = \left\{ \mathbf{x}^n : 2^{-n(\mathbf{H}(X) + \epsilon)} < p_{\mathbf{X}^n}(\mathbf{x}^n) < 2^{-n(\mathbf{H}(X) - \epsilon)} \right\}.$$

**Typical elements are approximately equiprobable in the strange sense above.**

**The complementary, atypical set of strings, satisfy**

$$\Pr[(T_\epsilon^n)^c] \leq \frac{\sigma_W^2}{n\epsilon^2}$$

**For any  $\epsilon, \delta > 0$ , large enough  $n$ ,  $\Pr[(T_\epsilon^n)^c] < \delta$ .**

**For all  $\mathbf{X}^n \in \mathbf{T}_\epsilon^n$ ,  $p_{\mathbf{X}^n}(\mathbf{X}^n) > 2^{-n[\mathbf{H}(\mathbf{X})+\epsilon]}$ .**

$$1 \geq \sum_{\mathbf{X}^n \in \mathbf{T}_\epsilon^n} p_{\mathbf{X}^n}(\mathbf{X}^n) > |\mathbf{T}_\epsilon^n| 2^{-n[\mathbf{H}(\mathbf{X})+\epsilon]}$$

$$|\mathbf{T}_\epsilon^n| < 2^{n[\mathbf{H}(\mathbf{X})+\epsilon]}$$

$$1 - \delta \leq \sum_{\mathbf{X}^n \in \mathbf{T}_\epsilon^n} p_{\mathbf{X}^n}(\mathbf{X}^n) < |\mathbf{T}_\epsilon^n| 2^{-n[\mathbf{H}(\mathbf{X})-\epsilon]}$$

$$|\mathbf{T}_\epsilon^n| > (1 - \delta) 2^{n[\mathbf{H}(\mathbf{X})-\epsilon]}$$

**Summary:**  $\Pr[(\mathbf{T}_\epsilon^n)^c] \approx 0$ ,  $|\mathbf{T}_\epsilon^n| \approx 2^{n\mathbf{H}(\mathbf{X})}$ ,

$p_{\mathbf{X}^n}(\mathbf{X}^n) \approx 2^{-n\mathbf{H}(\mathbf{X})}$  **for  $\mathbf{X}^n \in \mathbf{T}_\epsilon^n$ .**

## EXAMPLE

Consider binary DMS with  $\Pr[X=1] = p < 1/2$ .

$$H(X) = -p \log p - (1-p) \log(1-p)$$

The typical set  $T_\epsilon^n$  is the set of strings with about  $pn$  ones and  $(1-p)n$  zeros.

The probability of a typical string is about  $p^{pn}(1-p)^{(1-p)n} = 2^{-nH(X)}$ .

The number of  $n$ -strings with  $pn$  ones is  $\frac{n!}{(pn)!(n-pn)!}$

Note that there are  $2^n$  binary strings. Most of them are collectively very improbable.

The most probable strings have almost all zeros, but there aren't enough of them to matter.

## Fixed-to-fixed-length source codes

For any  $\epsilon, \delta > 0$ , and any large enough  $n$ , assign fixed length code word to each  $X^n \in T_\epsilon$ .

Since  $|T_\epsilon| < 2^{n[\mathbf{H}(X)+\epsilon]}$ ,  $\bar{L} \leq \mathbf{H}(X) + \epsilon + 1/n$ .

$$\Pr\{\text{failure}\} \leq \delta.$$

Conversely, take  $\bar{L} \leq \mathbf{H}(X) - 2\epsilon$ , and  $n$  large.

Since  $|T_\epsilon^n| > (1 - \delta)2^{n[\mathbf{H}(X)-\epsilon]}$ , most of typical set can not be assigned codewords.

$$\Pr\{\text{failure}\} > 1 - \delta - 2^{-\epsilon n} \rightarrow 1$$

## Kraft inequality for unique decodability

Suppose  $\{l_i\}$  are lengths of a uniquely decodable code and  $\sum_i 2^{-l_i} = b$ . We show that  $b > 1$  leads to contradiction. Choose DMS with  $p_i = (1/b)2^{-l_i}$ , i.e.,  $l_i = -\log(bp_i)$ .

$$\bar{L} = \sum_i p_i l_i = \mathbf{H}(X) - \log b$$

Consider string of  $n$  source letters. Concatenation of code words has length less than  $n[\mathbf{H}(X) - b/2]$  with high probability. Thus fixed length code of this length has low failure probability.

**Contradiction.**

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